

Simultaneous Testing of Mean Vector and Covariance Matrix with Three-step Monotone Missing Data

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Abstract

In this paper, we consider the problem of simultaneous testing of the mean vector and the covariance matrix in the case of a three-step monotone pattern. We provide the likelihood ratio test (LRT) statistic and propose a modified LRT statistic to improve the accuracy of the χ^2 approximation. This modified LRT statistic is derived by decomposing of the likelihood ratio (LR) using the coefficient of the modified LRT statistic with complete data. Finally, we investigate the asymptotic behavior of the upper percentiles of these test statistics via Monte Carlo simulation.

Key Words and Phrases. Likelihood ratio test, linear interpolation, maximum likelihood estimator, modified likelihood ratio test statistic.

1 Introduction

Examining the problem of missing data is an important aspect of statistical data analysis. Several authors have developed various statistical methods for dealing with missing data. In this paper, we consider the problem of simultaneous testing of the mean vector and the covariance matrix in the case of a three-step monotone pattern. For the simultaneous test, the LRT statistic and the modified LRT statistic with Bartlett correction in the case of complete data were discussed by Muirhead (1982) and Srivastava (2002). Further, an LRT statistic and a modified LRT statistic for two-step monotone missing data was proposed by Hao and Krishnamoorthy (2001) and Hosoya and Seo (2015). In particular, Hosoya and Seo (2015) presented a modified LRT statistic by decomposing the LR; this paper is an extension of the work presented by Hosoya and Seo (2015). Indeed, an LRT statistic and a modified LRT statistic for general k -step monotone missing data, which is obtained by correcting only a part of the missing data, was given by Yagi, Yamaguchi, and Seo (2016).

The remainder of this paper is organized as follows. In Section 2, we describe the MLEs of the mean vector and covariance matrix and its LRT statistic in the case of three-step monotone missing data. MLEs for general k -step monotone missing data were discussed by Jinadasa and Tracy

(1992) and Kanda and Fujikoshi (1998). Section 3 describes the modified LRT statistic in the case of complete data, which is used to derive the modified LRT statistic in Section 4. In Section 4, we decompose the LR and propose two modified LRT statistics. We also propose a modified LRT statistic using linear interpolation. In Section 5, using Monte Carlo simulation, we investigate the χ^2 approximation accuracy of the three modified LRT statistics proposed in Section 4. Finally, in Section 6, we state our conclusions.

2 Likelihood ratio with three-step monotone missing data

We suppose that the data is normally distributed as follows.

$$\begin{aligned} \mathbf{x}_1, \mathbf{x}_2, \dots, \mathbf{x}_{N_1} &\stackrel{i.i.d.}{\sim} N_p(\boldsymbol{\mu}, \boldsymbol{\Sigma}), \\ \mathbf{x}_{(12), N_1+1}, \mathbf{x}_{(12), N_1+2}, \dots, \mathbf{x}_{(12), N_1+N_2} &\stackrel{i.i.d.}{\sim} N_{p_1+p_2}(\boldsymbol{\mu}_{(12)}, \boldsymbol{\Sigma}_{(12)(12)}), \\ \mathbf{x}_{1, N_1+N_2+1}, \mathbf{x}_{1, N_1+N_2+2}, \dots, \mathbf{x}_{1N} &\stackrel{i.i.d.}{\sim} N_{p_1}(\boldsymbol{\mu}_1, \boldsymbol{\Sigma}_{11}), \end{aligned}$$

where

$$\begin{aligned} \boldsymbol{\mu} &= \begin{pmatrix} \boldsymbol{\mu}_1 \\ \boldsymbol{\mu}_2 \\ \boldsymbol{\mu}_3 \end{pmatrix} = \begin{pmatrix} \boldsymbol{\mu}_{(12)} \\ \boldsymbol{\mu}_3 \end{pmatrix}, \\ \boldsymbol{\Sigma} &= \left(\begin{array}{ccc|ccc} \boldsymbol{\Sigma}_{11} & \boldsymbol{\Sigma}_{12} & \boldsymbol{\Sigma}_{13} & & & \\ \boldsymbol{\Sigma}_{21} & \boldsymbol{\Sigma}_{22} & \boldsymbol{\Sigma}_{23} & & & \\ \boldsymbol{\Sigma}_{31} & \boldsymbol{\Sigma}_{32} & \boldsymbol{\Sigma}_{33} & & & \end{array} \right) = \left(\begin{array}{c|c} \boldsymbol{\Sigma}_{(12)(12)} & \boldsymbol{\Sigma}_{(12)3} \\ \hline \boldsymbol{\Sigma}_{3(12)} & \boldsymbol{\Sigma}_{33} \end{array} \right). \end{aligned}$$

We partition \mathbf{x}_j into a $p_1 \times 1$ random vector, a $p_2 \times 1$ random vector, and $p_3 \times 1$ random vector as $\mathbf{x}_j = (\mathbf{x}'_{1j}, \mathbf{x}'_{2j}, \mathbf{x}'_{3j})'$ ($j = 1, \dots, N_1$). Additionally, let $\mathbf{x}_{(12), j} = (\mathbf{x}'_{1j}, \mathbf{x}'_{2j})'$ ($j = N_1 + 1, \dots, N_1 + N_2$).

Such a dataset has three-step monotone missing data:

$$\left(\begin{array}{ccc|ccc} \overbrace{\mathbf{x}'_{11}} & \overbrace{\mathbf{x}'_{21}} & \overbrace{\mathbf{x}'_{31}} & & & \\ \vdots & \vdots & \vdots & & & \\ \mathbf{x}'_{1N_1} & \mathbf{x}'_{2N_1} & \mathbf{x}'_{3N_1} & & & \\ \mathbf{x}'_{1, N_1+1} & \mathbf{x}'_{2, N_1+1} & * \cdots * & & & \\ \vdots & \vdots & \vdots & & & \\ \mathbf{x}'_{1, N_1+N_2} & \mathbf{x}'_{2, N_1+N_2} & * \cdots * & & & \\ \mathbf{x}'_{1, N_1+N_2+1} & * \cdots * & * \cdots * & & & \\ \vdots & \vdots & \vdots & & & \\ \mathbf{x}'_{1N} & * \cdots * & * \cdots * & & & \end{array} \right),$$

where $N = N_1 + N_2 + N_3$, $p = p_1 + p_2 + p_3$ and “*” indicates a missing observation.

To derive the LRT statistic of the simultaneous testing of the mean vector and the covariance matrix in the case of a three-step monotone missing data pattern, we present their MLEs, which are given by

$$\begin{aligned}\hat{\boldsymbol{\mu}}_1 &= \frac{1}{N}(N_1\bar{\boldsymbol{x}}_{(1)1} + N_2\bar{\boldsymbol{x}}_{(2)1} + N_3\bar{\boldsymbol{x}}_{(3)}), \\ \hat{\boldsymbol{\mu}}_2 &= \frac{1}{N_1 + N_2}(N_1\bar{\boldsymbol{x}}_{(1)2} + N_2\bar{\boldsymbol{x}}_{(2)2}) - \boldsymbol{F}'\left(\frac{N_1\bar{\boldsymbol{x}}_{(1)1} + N_2\bar{\boldsymbol{x}}_{(2)1}}{N_1 + N_2} - \hat{\boldsymbol{\mu}}_1\right), \\ \hat{\boldsymbol{\mu}}_3 &= \bar{\boldsymbol{x}}_{(1)3} - \boldsymbol{G}'\begin{pmatrix} \bar{\boldsymbol{x}}_{(1)1} - \hat{\boldsymbol{\mu}}_1 \\ \bar{\boldsymbol{x}}_{(1)2} - \hat{\boldsymbol{\mu}}_2 \end{pmatrix}, \\ \hat{\boldsymbol{\Sigma}}_{11} &= \frac{1}{N}(\boldsymbol{W}_{(1)11} + \boldsymbol{W}_{(2)11} + \boldsymbol{W}_{(3)}), \quad \hat{\boldsymbol{\Sigma}}_{12} = \hat{\boldsymbol{\Sigma}}'_{21} = \hat{\boldsymbol{\Sigma}}_{11}\boldsymbol{F}, \\ \hat{\boldsymbol{\Sigma}}_{22} &= \frac{1}{N_1 + N_2}\boldsymbol{A}_{22\cdot 1} + \boldsymbol{F}'\hat{\boldsymbol{\Sigma}}_{11}\boldsymbol{F}, \quad \hat{\boldsymbol{\Sigma}}_{(12)3} = \hat{\boldsymbol{\Sigma}}'_{(3)12} = \hat{\boldsymbol{\Sigma}}_{(12)(12)}\boldsymbol{G}, \\ \hat{\boldsymbol{\Sigma}}_{33} &= \frac{1}{N_1}\boldsymbol{W}_{(1)33\cdot 12} + \boldsymbol{G}'\hat{\boldsymbol{\Sigma}}_{(12)(12)}\boldsymbol{G},\end{aligned}$$

where

$$\begin{aligned}\bar{\boldsymbol{x}}_{(1)} &= \begin{pmatrix} \bar{\boldsymbol{x}}_{(1)1} \\ \bar{\boldsymbol{x}}_{(1)2} \\ \bar{\boldsymbol{x}}_{(1)3} \end{pmatrix}, \quad \bar{\boldsymbol{x}}_{(1)1} = \frac{1}{N_1}\sum_{j=1}^{N_1}\boldsymbol{x}_{1j}, \quad \bar{\boldsymbol{x}}_{(1)2} = \frac{1}{N_1}\sum_{j=1}^{N_1}\boldsymbol{x}_{2j}, \quad \bar{\boldsymbol{x}}_{(1)3} = \frac{1}{N_1}\sum_{j=1}^{N_1}\boldsymbol{x}_{3j}, \\ \bar{\boldsymbol{x}}_{(2)} &= \begin{pmatrix} \bar{\boldsymbol{x}}_{(2)1} \\ \bar{\boldsymbol{x}}_{(2)2} \end{pmatrix}, \quad \bar{\boldsymbol{x}}_{(2)1} = \frac{1}{N_2}\sum_{j=N_1+1}^{N_1+N_2}\boldsymbol{x}_{1j}, \quad \bar{\boldsymbol{x}}_{(2)2} = \frac{1}{N_2}\sum_{j=N_1+1}^{N_1+N_2}\boldsymbol{x}_{2j}, \quad \bar{\boldsymbol{x}}_{(3)} = \frac{1}{N_3}\sum_{j=N_1+N_2+1}^N\boldsymbol{x}_{1j},\end{aligned}$$

and

$$\begin{aligned}\boldsymbol{F} &= (\boldsymbol{W}_{(1)11} + \boldsymbol{W}_{(2)11})^{-1}(\boldsymbol{W}_{(1)12} + \boldsymbol{W}_{(2)12}), \quad \boldsymbol{G} = (\boldsymbol{W}_{(1),(12)(12)})^{-1}\boldsymbol{W}_{(1),(12)3}, \\ \boldsymbol{W}_{(1)} &= \sum_{j=1}^{N_1}(\boldsymbol{x}_j - \bar{\boldsymbol{x}}_{(1)})(\boldsymbol{x}_j - \bar{\boldsymbol{x}}_{(1)})' = \begin{pmatrix} \boldsymbol{W}_{(1)11} & \boldsymbol{W}_{(1)12} & \boldsymbol{W}_{(1)13} \\ \boldsymbol{W}_{(1)21} & \boldsymbol{W}_{(1)22} & \boldsymbol{W}_{(1)23} \\ \boldsymbol{W}_{(1)31} & \boldsymbol{W}_{(1)32} & \boldsymbol{W}_{(1)33} \end{pmatrix} = \begin{pmatrix} \boldsymbol{W}_{(1),(12)(12)} & \boldsymbol{W}_{(1),(12)3} \\ \boldsymbol{W}_{(1),3(12)} & \boldsymbol{W}_{(1)33} \end{pmatrix}, \\ \boldsymbol{W}_{(2)} &= \sum_{j=N_1+1}^{N_1+N_2} \begin{pmatrix} \boldsymbol{x}_{1j} - \bar{\boldsymbol{x}}_{(2)1} \\ \boldsymbol{x}_{2j} - \bar{\boldsymbol{x}}_{(2)2} \end{pmatrix} \begin{pmatrix} \boldsymbol{x}_{1j} - \bar{\boldsymbol{x}}_{(2)1} \\ \boldsymbol{x}_{2j} - \bar{\boldsymbol{x}}_{(2)2} \end{pmatrix}' + \frac{N_1N_2}{N_1 + N_2} \begin{pmatrix} \bar{\boldsymbol{x}}_{(1)1} - \bar{\boldsymbol{x}}_{(2)1} \\ \bar{\boldsymbol{x}}_{(1)2} - \bar{\boldsymbol{x}}_{(2)2} \end{pmatrix} \begin{pmatrix} \bar{\boldsymbol{x}}_{(1)1} - \bar{\boldsymbol{x}}_{(2)1} \\ \bar{\boldsymbol{x}}_{(1)2} - \bar{\boldsymbol{x}}_{(2)2} \end{pmatrix}' \\ &= \begin{pmatrix} \boldsymbol{W}_{(2)11} & \boldsymbol{W}_{(2)12} \\ \boldsymbol{W}_{(2)21} & \boldsymbol{W}_{(2)22} \end{pmatrix}, \\ \boldsymbol{W}_{(3)} &= \sum_{j=N_1+N_2+1}^N (\boldsymbol{x}_{1j} - \bar{\boldsymbol{x}}_{(3)})(\boldsymbol{x}_{1j} - \bar{\boldsymbol{x}}_{(3)})' + \frac{(N_1 + N_2)N_3}{N} \left(\bar{\boldsymbol{x}}_{(3)} - \frac{1}{N_1 + N_2}(N_1\bar{\boldsymbol{x}}_{(1)1} + N_2\bar{\boldsymbol{x}}_{(2)1}) \right) \\ &\quad \times \left(\bar{\boldsymbol{x}}_{(3)} - \frac{1}{N_1 + N_2}(N_1\bar{\boldsymbol{x}}_{(1)1} + N_2\bar{\boldsymbol{x}}_{(2)1}) \right)',\end{aligned}$$

$$\boldsymbol{W}_{(1)33\cdot 12} = \boldsymbol{W}_{(1)33} - \boldsymbol{W}_{(1),3(12)}\boldsymbol{W}_{(1),(12)(12)}^{-1}\boldsymbol{W}_{(1),(12)3},$$

$$\boldsymbol{A} = \boldsymbol{W}_{(1),(12)(12)} + \boldsymbol{W}_{(2)}, \quad \boldsymbol{A}_{22\cdot 1} = \boldsymbol{A}_{22} - \boldsymbol{A}_{21}\boldsymbol{A}_{11}^{-1}\boldsymbol{A}_{12}.$$

The results obtained are the same as those obtained in the work of Kanda and Fujikoshi (1998). For details, see Appendix. Now, consider the following hypothesis test when the dataset has a three-step monotone pattern.

$$H_0 : \boldsymbol{\mu} = \boldsymbol{\mu}_0, \boldsymbol{\Sigma} = \boldsymbol{\Sigma}_0 \text{ vs. } H_1 : \text{not } H_0,$$

where $\boldsymbol{\mu}_0$ is a known vector and $\boldsymbol{\Sigma}_0$ is a known matrix. Without loss of generality, we can assume that $\boldsymbol{\mu} = \mathbf{0}$ and $\boldsymbol{\Sigma} = \mathbf{I}_p$. Then, the LR of the hypothesis test can be given by

$$\begin{aligned} \lambda = & \left(\frac{e}{N}\right)^{\frac{1}{2}Np_1} |\mathbf{W}_{(1)11} + \mathbf{W}_{(2)11} + \mathbf{W}_{(3)}|^{-\frac{N}{2}} \left(\frac{e}{N_1 + N_2}\right)^{\frac{1}{2}(N_1 + N_2)p_2} |\mathbf{A}_{22 \cdot 1}|^{-\frac{N_1 + N_2}{2}} \left(\frac{e}{N_1}\right)^{\frac{1}{2}N_1p_3} |\mathbf{W}_{(1)33 \cdot 12}|^{-\frac{N_1}{2}} \\ & \times \text{etr} \left(-\frac{1}{2} (\mathbf{W}_{(1)11} + \mathbf{W}_{(2)11} + \mathbf{W}_{(3)}) + \frac{1}{N} (N_1 \bar{\mathbf{x}}_{(1)1} + N_2 \bar{\mathbf{x}}_{(2)1} + N_3 \bar{\mathbf{x}}_{(3)}) (N_1 \bar{\mathbf{x}}_{(1)1} + N_2 \bar{\mathbf{x}}_{(2)1} + N_3 \bar{\mathbf{x}}_{(3)})' \right) \\ & \times \text{etr} \left(-\frac{1}{2} (\mathbf{A}_{22 \cdot 1} + \mathbf{A}_{21} \mathbf{A}_{11}^{-1} \mathbf{A}_{12} + \frac{1}{N_1 + N_2} (N_1 \bar{\mathbf{x}}_{(1)2} + N_2 \bar{\mathbf{x}}_{(2)2}) (N_1 \bar{\mathbf{x}}_{(1)2} + N_2 \bar{\mathbf{x}}_{(2)2})') \right) \\ & \times \text{etr} \left(-\frac{1}{2} (\mathbf{W}_{(1)33 \cdot 12} + \mathbf{W}_{(1),3(12)} \mathbf{W}_{(1),(12)(12)}^{-1} \mathbf{W}_{(1),(12)3} + N_1 \bar{\mathbf{x}}_{(1)3} \bar{\mathbf{x}}_{(1)3}') \right). \end{aligned}$$

This LR is essentially the same that obtained by Yagi, Yamaguchi, and Seo (2016). Thus, we obtain the test statistic $-2 \log \lambda$. For example, Table 1 presents the simulated values of the upper 100α percentiles of $-2 \log \lambda$ and Type I error rate, $\alpha_1 = \Pr\{-2 \log \lambda > \chi_f^2(\alpha)\}$ for the three-step monotone missing data case, where $\chi_f^2(\alpha)$ is the upper percentile of the χ^2 distribution with f degrees of freedom and $f = p(p + 3)/2$.

Table 1: The upper percentile of $-2 \log \lambda$ and type I error rates when $(p_1, p_2, p_3) = (3, 3, 3)$.

Sample Size			Upper Percentile	Type I Error Rate
N_1	N_2	N_3	$-2 \log \lambda$	α_1
<u>$\alpha = 0.05$</u>				
10	10	10	155.129	0.907
20	10	10	87.775	0.282
40	10	10	78.955	0.127
80	10	10	75.419	0.082
200	10	10	73.431	0.061
<u>$\alpha = 0.01$</u>				
10	10	10	187.825	0.804
20	10	10	99.009	0.115
40	10	10	88.653	0.036
80	10	10	84.693	0.019
200	10	10	82.514	0.013

Note. $\chi_f^2(0.05) = 72.15$, $\chi_f^2(0.01) = 81.07$, $f = 54$.

As presented in Table 1, the accuracy of the χ^2 approximation in this case is not desirable when

the sample size is not large; therefore, a modified LRT statistic is needed to improve the accuracy of the χ^2 approximation.

3 LRT statistic for complete data

We consider the LRT statistic and the modified LRT statistic with Bartlett correction in the case of complete data. These results are used in the next section. We first consider a simultaneous test for complete data as follows.

$$H_{01} : \boldsymbol{\mu} = \mathbf{0}, \boldsymbol{\Sigma} = \mathbf{I} \text{ vs. } H_{11} : \text{not } H_{01}$$

In this case, the LR can be expressed as follows. Let $\mathbf{x}_1, \mathbf{x}_2, \dots, \mathbf{x}_N$ be independently distributed as $N_p(\boldsymbol{\mu}, \boldsymbol{\Sigma})$, and let λ_S be the LR for the complete data. Then, the LR is given by

$$\lambda_S = e^{\frac{Np}{2}} \left| \frac{1}{N} \mathbf{W} \right|^{\frac{N}{2}} \text{etr} \left(-\frac{1}{2} \mathbf{W} \right) \exp \left(-\frac{1}{2} N \bar{\mathbf{x}}' \bar{\mathbf{x}} \right),$$

where

$$\mathbf{W} = \sum_{j=1}^N (\mathbf{x}_j - \bar{\mathbf{x}})(\mathbf{x}_j - \bar{\mathbf{x}})', \quad \bar{\mathbf{x}} = \frac{1}{N} \sum_{j=1}^N \mathbf{x}_j.$$

Furthermore, the modified LRT statistic with Bartlett correction can be given by $-2\rho_1 \log \lambda_S$ (Muirhead (1982, p370)), where

$$\rho_1 = 1 - \frac{2p^2 + 9p + 11}{6N(p + 3)}.$$

Next, we consider a covariance test for complete data as follows.

$$H_{02} : \boldsymbol{\Sigma} = \mathbf{I} \text{ vs. } H_{12} : \text{not } H_{02}$$

In this case, the LR, which is an unbiased test, can be expressed as follows.

$$\lambda_V = e^{\frac{(N-1)p}{2}} \left| \frac{1}{N-1} \mathbf{W} \right|^{\frac{N-1}{2}} \text{etr} \left(-\frac{1}{2} \mathbf{W} \right).$$

The modified LRT statistic with Bartlett correction $-2\rho_2 \log \lambda_V$ was provided by Muirhead (1982, p359), where

$$\rho_2 = 1 - \frac{2p^2 + 3p - 1}{6(N-1)(p+1)}.$$

4 Modified LRT statistic

We now decompose the LR to derive the modified LRT statistic. Let

$$\begin{aligned}
\omega_1 &= \exp \left(-\frac{1}{2N} (N_1 \bar{\mathbf{x}}_{(1)1} + N_2 \bar{\mathbf{x}}_{(2)1} + N_3 \bar{\mathbf{x}}_{(3)})' (N_1 \bar{\mathbf{x}}_{(1)1} + N_2 \bar{\mathbf{x}}_{(2)1} + N_3 \bar{\mathbf{x}}_{(3)}) \right), \\
\omega_2 &= \exp \left(-\frac{1}{2(N_1 + N_2)} (N_1 \bar{\mathbf{x}}_{(1)2} + N_2 \bar{\mathbf{x}}_{(2)2})' (N_1 \bar{\mathbf{x}}_{(1)2} + N_2 \bar{\mathbf{x}}_{(2)2}) \right), \\
\omega_3 &= \exp \left(-\frac{N_1}{2} \bar{\mathbf{x}}'_{(1)3} \bar{\mathbf{x}}_{(1)3} \right), \\
\omega_4 &= \left(\frac{e}{N} \right)^{\frac{1}{2} N p_1} |\mathbf{W}_{(1)11} + \mathbf{W}_{(2)11} + \mathbf{W}_{(3)}|^{\frac{N}{2}} \text{etr} \left(-\frac{1}{2} (\mathbf{W}_{(1)11} + \mathbf{W}_{(2)11} + \mathbf{W}_{(3)}) \right), \\
\omega_5 &= \left(\frac{e}{N_1 + N_2} \right)^{\frac{1}{2} (N_1 + N_2) p_2} |\mathbf{A}_{22 \cdot 1}|^{\frac{N_1 + N_2}{2}} \text{etr} \left(-\frac{1}{2} \mathbf{A}_{22 \cdot 1} \right), \\
\omega_6 &= \left(\frac{e}{N_1} \right)^{\frac{1}{2} N_1 p_3} |\mathbf{W}_{(1)33 \cdot 12}|^{\frac{N_1}{2}} \text{etr} \left(-\frac{1}{2} \mathbf{W}_{(1)33 \cdot 12} \right), \\
\omega_7 &= \text{etr} \left(-\frac{1}{2} \mathbf{A}_{21} \mathbf{A}_{11}^{-1} \mathbf{A}_{12} \right) \text{etr} \left(-\frac{1}{2} \mathbf{W}_{(1),3(12)} \mathbf{W}_{(1),(12)(12)}^{-1} \mathbf{W}_{(1),(12)3} \right).
\end{aligned}$$

Therefore,

$$\lambda = \prod_{i=1}^7 \omega_i.$$

Then, $\omega_1 \omega_4, \omega_2 \omega_5, \omega_3 \omega_6$ are of the form of LR for H_{01} under non-missing normality. Hence, we can obtain the modified LRT statistic, $-2\rho_{14} \log \omega_1 \omega_4, -2\rho_{25} \log \omega_2 \omega_5, -2\rho_{36} \log \omega_3 \omega_6$, where

$$\rho_{14} = 1 - \frac{2p_1^2 + 9p_1 + 11}{6N(p_1 + 3)}, \quad \rho_{25} = 1 - \frac{2p_2^2 + 9p_2 + 11}{6(N_1 + N_2)(p_2 + 3)}, \quad \rho_{36} = 1 - \frac{2p_3^2 + 9p_3 + 11}{6N_1(p_3 + 3)}.$$

Thus, we propose a new modified LRT statistic given by $-2 \log \tau$, where

$$\tau = (\omega_1 \omega_4)^{\rho_{14}} (\omega_2 \omega_5)^{\rho_{25}} (\omega_3 \omega_6)^{\rho_{36}} \omega_7.$$

In addition, we denote

$$\begin{aligned}
\omega_4^* &= \left(\frac{e}{n} \right)^{\frac{1}{2} n p_1} |\mathbf{W}_{(1)11} + \mathbf{W}_{(2)11} + \mathbf{W}_{(3)}|^{\frac{n}{2}} \text{etr} \left(-\frac{1}{2} (\mathbf{W}_{(1)11} + \mathbf{W}_{(2)11} + \mathbf{W}_{(3)}) \right), \\
\omega_5^* &= \left(\frac{e}{n_1 + n_2} \right)^{\frac{1}{2} (n_1 + n_2) p_2} |\mathbf{A}_{22 \cdot 1}|^{\frac{n_1 + n_2}{2}} \text{etr} \left(-\frac{1}{2} \mathbf{A}_{22 \cdot 1} \right), \\
\omega_6^* &= \left(\frac{e}{n_1} \right)^{\frac{1}{2} n_1 p_3} |\mathbf{W}_{(1)33 \cdot 12}|^{\frac{n_1}{2}} \text{etr} \left(-\frac{1}{2} \mathbf{W}_{(1)33 \cdot 12} \right),
\end{aligned}$$

where

$$n = N - 1, \quad n_1 = N_1 - (p_1 + p_2) - 1, \quad n_1 + n_2 = N_1 + N_2 - p_1 - 1.$$

Then, since $\omega_4^*, \omega_5^*, \omega_6^*$ are of the form of LR for H_{02} under non-missing normality, we can propose the modified LRT statistic as $-2 \log \phi$, where

$$\phi = \omega_1 \omega_2 \omega_3 (\omega_4^*)^{\rho_4^*} (\omega_5^*)^{\rho_5^*} (\omega_6^*)^{\rho_6^*} \omega_7$$

and

$$\rho_4^* = 1 - \frac{2p_1^2 + 3p_1 - 1}{6n(p_1 + 1)}, \quad \rho_5^* = 1 - \frac{2p_2^2 + 3p_2 - 1}{6(n_1 + n_2)(p_2 + 1)}, \quad \rho_6^* = 1 - \frac{2p_3^2 + 3p_3 - 1}{6n_1(p_3 + 1)}.$$

Now, we propose the modified LRT statistic $-2\rho_m \log \lambda$ via linear interpolation, where

$$\rho_m = \left\{ 1 - \frac{(p_1 + p_2)N_2 + p_1N_3}{p(N_2 + N_3)} \right\} \rho_{N_1} + \frac{(p_1 + p_2)N_2 + p_1N_3}{p(N_2 + N_3)} \rho_N,$$

and

$$\rho_{N_1} = 1 - \frac{2p^2 + 9p + 11}{6N_1(p + 3)}, \quad \rho_N = 1 - \frac{2p^2 + 9p + 11}{6N(p + 3)}.$$

5 Simulation studies

We evaluate the accuracy and the asymptotic behaviors of the χ^2 approximations via Monte Carlo simulation (10^6 runs). Now let

$$\begin{aligned} \alpha_1 &= \Pr\{-2 \log \lambda > \chi_f^2(\alpha)\}, \\ \alpha_{\rho_m} &= \Pr\{-2\rho_m \log \lambda > \chi_f^2(\alpha)\}, \\ \alpha_\tau &= \Pr\{-2 \log \tau > \chi_f^2(\alpha)\}, \\ \alpha_\phi &= \Pr\{-2 \log \phi > \chi_f^2(\alpha)\}, \end{aligned}$$

where $\chi_f^2(\alpha)$ is the upper percentile of the χ^2 distribution with f degrees of freedom.

In Tables 2, 4, 6, and 8, we provide the simulated upper 100α percentiles and the actual type I error rates for the upper percentiles $\alpha_1, \alpha_{\rho_m}, \alpha_\tau$, and α_ϕ of $-2 \log \lambda, -2\rho_m \log \lambda, -2 \log \tau$ and $-2 \log \phi$; $\alpha = 0.05, 0.01$; and for the following cases (Case I),

$$(N_1, N_2, N_3) = \begin{cases} (m, 10, 10), \\ (m, 20, 20), & m = 10, 20, 30, 40, 80, 200, 400, \\ (m, 50, 50), \end{cases}$$

where (p_1, p_2, p_3) in Tables 2, 4, 6, and 8 are $(3, 3, 3), (4, 4, 4), (5, 5, 5), (6, 6, 6)$, respectively.

Similarly, Tables 3, 5, 7, and 9 show the results for the following cases (Case II),

$$(N_1, N_2, N_3) = \begin{cases} (m, m/2, m/2), \\ (m, m, m), & m = 10, 20, 30, 40, 80, 200, 400, \\ (m, 2m, 2m), \end{cases}$$

where (p_1, p_2, p_3) in Tables 3, 5, 7, and 9 are $(3, 3, 3), (4, 4, 4), (5, 5, 5), (6, 6, 6)$ respectively.

It may be noted from above mentioned Tables that the simulated values are closer to the upper percentile of the χ^2 distribution when the sample size increases. In addition, it can be seen that the upper percentile of $-2 \log \phi$ is considerably better than that of $-2 \log \lambda$ even for small sample sizes; while it is not as good as $-2 \log \phi$, the upper percentile of $-2\rho_m \log \lambda$ is good.

Table 2: The simulated values for $-2 \log \lambda$, $-2\rho_m \log \lambda$, $-2 \log \tau$, and $-2 \log \phi$ and type I error rates α_1 , α_{ρ_m} , α_τ , and α_ϕ when $(p_1, p_2, p_3) = (3, 3, 3)$ for Case I.

Sample Size			Upper Percentile				Type I Error Rate			
N_1	N_2	N_3	$-2 \log \lambda$	$-2\rho_m \log \lambda$	$-2 \log \tau$	$-2 \log \phi$	α_1	α_{ρ_m}	α_τ	α_ϕ
$\alpha = 0.05$										
10	10	10	155.129	118.645	137.865	74.060	0.907	0.597	0.833	0.067
20	10	10	87.775	76.163	84.460	72.172	0.282	0.090	0.220	0.050
40	10	10	78.955	73.152	77.524	72.168	0.127	0.059	0.108	0.050
80	10	10	75.419	72.426	74.720	72.177	0.082	0.052	0.074	0.050
200	10	10	73.431	72.195	73.151	72.158	0.061	0.050	0.059	0.050
400	10	10	72.767	72.140	72.625	72.140	0.055	0.050	0.054	0.050
10	20	20	153.232	120.798	136.674	73.968	0.895	0.619	0.822	0.066
20	20	20	86.852	76.639	83.930	72.184	0.263	0.095	0.209	0.050
40	20	20	78.565	73.368	77.286	72.154	0.122	0.060	0.105	0.050
80	20	20	75.299	72.532	74.652	72.165	0.080	0.053	0.073	0.050
200	20	20	73.410	72.223	73.143	72.142	0.061	0.051	0.058	0.050
400	20	20	72.812	72.199	72.677	72.183	0.056	0.050	0.055	0.050
10	50	50	151.864	122.642	135.795	74.019	0.883	0.634	0.810	0.067
20	50	50	85.844	77.011	83.319	72.207	0.242	0.100	0.197	0.050
40	50	50	78.011	73.588	76.975	72.158	0.115	0.063	0.101	0.050
80	50	50	75.046	72.656	74.505	72.153	0.078	0.054	0.072	0.050
200	50	50	73.371	72.293	73.128	72.145	0.061	0.051	0.058	0.050
400	50	50	72.781	72.203	72.653	72.177	0.055	0.050	0.054	0.050
$\alpha = 0.01$										
10	10	10	187.825	143.652	165.574	83.493	0.804	0.417	0.686	0.016
20	10	10	99.009	85.911	95.164	81.092	0.115	0.023	0.079	0.010
40	10	10	88.653	82.137	87.035	80.994	0.036	0.012	0.028	0.010
80	10	10	84.693	81.332	83.905	81.075	0.019	0.011	0.017	0.010
200	10	10	82.514	81.125	82.205	81.048	0.013	0.010	0.012	0.010
400	10	10	81.873	81.168	81.714	81.113	0.012	0.010	0.011	0.010
10	20	20	186.218	146.802	164.463	83.488	0.784	0.439	0.669	0.015
20	20	20	97.770	86.273	94.351	81.019	0.103	0.024	0.073	0.010
40	20	20	88.270	82.432	86.831	81.049	0.034	0.013	0.027	0.010
80	20	20	84.641	81.531	83.916	81.104	0.019	0.011	0.017	0.010
200	20	20	82.518	81.184	82.222	81.073	0.013	0.010	0.012	0.010
400	20	20	81.776	81.087	81.623	81.073	0.012	0.010	0.011	0.010
10	50	50	184.963	149.372	163.853	83.441	0.766	0.456	0.654	0.015
20	50	50	96.671	86.724	93.675	81.061	0.092	0.026	0.068	0.010
40	50	50	87.633	82.665	86.433	81.014	0.031	0.013	0.026	0.010
80	50	50	84.325	81.639	83.709	81.092	0.018	0.011	0.016	0.010
200	50	50	82.429	81.217	82.159	81.058	0.013	0.010	0.012	0.010
400	50	50	81.761	81.112	81.620	81.068	0.011	0.010	0.011	0.010

Note. The closest to α in the values $\alpha_1, \alpha_{\rho_m}, \alpha_\tau$, and α_ϕ of each row is in bold. $\chi_f^2(0.05) = 72.15$, $\chi_f^2(0.01) = 81.07$, $f = 54$.

Table 3: The simulated values for $-2 \log \lambda$, $-2\rho_m \log \lambda$, $-2 \log \tau$, and $-2 \log \phi$ and type I error rates α_1 , α_{ρ_m} , α_τ , and α_ϕ when $(p_1, p_2, p_3) = (3, 3, 3)$ for Case II.

Sample Size			Upper Percentile				Type I Error Rate			
N_1	N_2	N_3	$-2 \log \lambda$	$-2\rho_m \log \lambda$	$-2 \log \tau$	$-2 \log \phi$	α_1	α_{ρ_m}	α_τ	α_ϕ
$\alpha = 0.05$										
10	5	5	55.845	45.219	51.759	40.161	0.345	0.119	0.257	0.051
20	10	10	87.775	76.163	84.460	72.172	0.282	0.090	0.220	0.050
40	20	20	42.534	40.510	41.958	40.131	0.080	0.054	0.072	0.050
80	40	40	41.308	40.326	41.037	40.142	0.064	0.052	0.061	0.050
200	100	100	40.561	40.175	40.462	40.111	0.055	0.051	0.054	0.050
400	200	200	72.714	72.233	72.611	72.147	0.055	0.051	0.054	0.050
10	10	10	54.991	45.690	51.315	40.164	0.324	0.127	0.245	0.051
20	20	20	86.852	76.639	83.930	72.184	0.263	0.095	0.209	0.050
40	40	40	42.361	40.570	41.864	40.134	0.078	0.055	0.071	0.050
80	80	80	41.195	40.324	40.969	40.126	0.062	0.052	0.060	0.050
200	200	200	40.509	40.166	40.422	40.085	0.054	0.051	0.053	0.050
400	400	400	72.687	72.259	72.598	72.156	0.055	0.051	0.054	0.050
10	20	20	54.341	46.069	50.979	40.194	0.306	0.132	0.236	0.051
20	40	40	44.857	41.443	43.822	40.097	0.118	0.066	0.100	0.050
40	80	80	42.199	40.593	41.774	40.143	0.076	0.055	0.070	0.050
80	160	160	41.065	40.283	40.874	40.088	0.061	0.052	0.059	0.050
200	400	400	40.490	40.182	40.416	40.113	0.054	0.051	0.053	0.050
400	800	800	72.676	72.292	72.601	72.198	0.054	0.051	0.054	0.050
$\alpha = 0.01$										
10	5	5	190.179	139.861	167.021	83.493	0.828	0.385	0.706	0.015
20	10	10	99.009	85.911	95.164	81.092	0.115	0.023	0.079	0.010
40	20	20	88.389	82.543	86.929	81.141	0.034	0.013	0.027	0.010
80	40	40	84.349	81.560	83.702	81.078	0.018	0.011	0.016	0.010
200	100	100	82.312	81.223	82.081	81.080	0.013	0.010	0.012	0.010
400	200	200	81.762	81.221	81.644	81.138	0.011	0.010	0.011	0.010
10	10	10	188.082	143.848	165.818	83.476	0.804	0.417	0.686	0.016
20	20	20	97.770	86.273	94.351	81.019	0.103	0.024	0.073	0.010
40	40	40	87.906	82.737	86.645	81.122	0.032	0.014	0.026	0.010
80	80	80	84.169	81.695	83.605	81.092	0.018	0.011	0.016	0.010
200	200	200	82.283	81.316	82.084	81.101	0.013	0.010	0.012	0.010
400	400	400	81.592	81.112	81.492	81.002	0.011	0.010	0.011	0.010
10	20	20	185.957	146.596	164.298	83.484	0.783	0.438	0.668	0.015
20	40	40	96.935	86.676	93.860	81.140	0.095	0.026	0.069	0.010
40	80	80	87.339	82.717	86.232	80.993	0.030	0.014	0.025	0.010
80	160	160	84.036	81.813	83.565	81.047	0.017	0.012	0.016	0.010
200	400	400	82.216	81.346	82.052	81.089	0.012	0.011	0.012	0.010
400	800	800	81.654	81.222	81.570	81.100	0.011	0.010	0.011	0.010

Note. The closest to α in the values α_1 , α_{ρ_m} , α_τ , and α_ϕ of each row is in bold. $\chi_f^2(0.05) = 72.15$, $\chi_f^2(0.01) = 81.07$, $f = 54$.

Table 4: The simulated values for $-2 \log \lambda$, $-2\rho_m \log \lambda$, $-2 \log \tau$, and $-2 \log \phi$ and type I error rates α_1 , α_{ρ_m} , α_τ , and α_ϕ when $(p_1, p_2, p_3) = (4, 4, 4)$ for Case I.

Sample Size			Upper Percentile				Type I Error Rate			
N_1	N_2	N_3	$-2 \log \lambda$	$-2\rho_m \log \lambda$	$-2 \log \tau$	$-2 \log \phi$	α_1	α_{ρ_m}	α_τ	α_ϕ
$\alpha = 0.05$										
20	10	10	149.889	124.470	142.692	113.298	0.591	0.160	0.473	0.051
30	10	10	133.666	117.547	129.734	113.243	0.313	0.084	0.246	0.051
40	10	10	127.521	115.507	124.780	113.127	0.211	0.067	0.171	0.050
80	10	10	119.858	113.760	118.578	113.136	0.109	0.054	0.096	0.050
200	10	10	115.712	113.215	115.213	113.103	0.069	0.050	0.065	0.050
400	10	10	114.393	113.131	114.144	113.117	0.059	0.050	0.057	0.050
20	20	20	147.970	125.665	141.494	113.340	0.557	0.176	0.451	0.051
30	20	20	132.423	118.165	128.942	113.163	0.292	0.091	0.235	0.050
40	20	20	126.793	116.042	124.338	113.132	0.200	0.071	0.165	0.050
80	20	20	119.598	113.964	118.419	113.161	0.106	0.056	0.094	0.050
200	20	20	115.671	113.273	115.189	113.111	0.069	0.051	0.065	0.050
400	20	20	114.396	113.162	114.153	113.115	0.059	0.050	0.057	0.050
20	50	50	145.872	126.631	140.151	113.292	0.518	0.189	0.425	0.051
30	50	50	130.985	118.834	128.030	113.232	0.266	0.097	0.219	0.051
40	50	50	125.690	116.555	123.648	113.168	0.184	0.076	0.156	0.050
80	50	50	119.099	114.236	118.099	113.111	0.102	0.058	0.091	0.050
200	50	50	115.644	113.465	115.212	113.186	0.068	0.052	0.065	0.050
400	50	50	114.438	113.273	114.208	113.189	0.059	0.051	0.057	0.050
$\alpha = 0.01$										
20	10	10	165.180	137.168	157.031	124.325	0.356	0.052	0.248	0.010
30	10	10	146.744	129.048	142.334	124.223	0.131	0.021	0.092	0.010
40	10	10	140.003	126.813	136.940	124.160	0.074	0.015	0.055	0.010
80	10	10	131.473	124.784	130.084	124.103	0.029	0.011	0.024	0.010
200	10	10	126.859	124.121	126.316	123.989	0.015	0.010	0.014	0.010
400	10	10	125.529	124.143	125.250	124.124	0.012	0.010	0.012	0.010
20	20	20	163.201	138.600	155.800	124.362	0.323	0.059	0.229	0.010
30	20	20	145.442	129.782	141.534	124.107	0.118	0.023	0.085	0.010
40	20	20	139.231	127.425	136.508	124.193	0.068	0.016	0.051	0.010
80	20	20	131.310	125.125	130.024	124.203	0.028	0.012	0.024	0.010
200	20	20	126.788	124.160	126.275	123.965	0.015	0.010	0.014	0.010
400	20	20	125.588	124.233	125.319	124.207	0.013	0.010	0.012	0.010
20	50	50	160.721	139.522	154.110	124.237	0.288	0.065	0.209	0.010
30	50	50	143.756	130.421	140.409	124.017	0.102	0.025	0.077	0.010
40	50	50	137.921	127.897	135.635	124.119	0.060	0.018	0.047	0.010
80	50	50	130.681	125.345	129.576	124.154	0.026	0.012	0.022	0.010
200	50	50	126.745	124.357	126.270	124.110	0.015	0.010	0.014	0.010
400	50	50	125.467	124.191	125.214	124.098	0.012	0.010	0.012	0.010

Note. The closest to α in the values $\alpha_1, \alpha_{\rho_m}, \alpha_\tau$, and α_ϕ of each row is in bold. $\chi_f^2(0.05) = 113.15$, $\chi_f^2(0.01) = 124.12$, $f = 90$.

Table 5: The simulated values for $-2 \log \lambda$, $-2\rho_m \log \lambda$, $-2 \log \tau$, and $-2 \log \phi$ and type I error rates $\alpha_1, \alpha_{\rho_m}, \alpha_\tau$, and α_ϕ when $(p_1, p_2, p_3) = (4, 4, 4)$ for Case II.

Sample Size			Upper Percentile				Type I Error Rate			
N_1	N_2	N_3	$-2 \log \lambda$	$-2\rho_m \log \lambda$	$-2 \log \tau$	$-2 \log \phi$	α_1	α_{ρ_m}	α_τ	α_ϕ
$\alpha = 0.05$										
20	10	10	149.889	124.470	142.692	113.298	0.591	0.160	0.473	0.051
30	15	15	132.978	117.944	129.307	113.214	0.302	0.089	0.240	0.050
40	20	20	126.775	116.025	124.320	113.147	0.200	0.071	0.165	0.050
80	40	40	119.344	114.285	118.295	113.196	0.103	0.058	0.092	0.050
200	100	100	115.375	113.418	114.994	113.082	0.066	0.052	0.063	0.050
400	200	200	114.317	113.348	114.128	113.156	0.058	0.051	0.057	0.050
20	20	20	147.970	125.665	141.494	113.340	0.557	0.176	0.451	0.051
30	30	30	131.808	118.562	128.583	113.223	0.281	0.095	0.229	0.050
40	40	40	125.957	116.463	123.820	113.164	0.188	0.076	0.158	0.050
80	80	80	118.944	114.462	118.044	113.206	0.099	0.059	0.090	0.050
200	200	200	115.324	113.586	114.999	113.146	0.066	0.053	0.063	0.050
400	400	400	114.155	113.295	113.998	113.074	0.057	0.051	0.056	0.050
20	40	40	146.236	126.397	140.373	113.279	0.526	0.186	0.430	0.051
30	60	60	130.756	118.930	127.897	113.228	0.262	0.098	0.216	0.051
40	80	80	125.157	116.667	123.279	113.120	0.176	0.077	0.150	0.050
80	160	160	118.479	114.460	117.718	113.155	0.094	0.059	0.087	0.050
200	400	400	115.160	113.598	114.889	113.144	0.064	0.053	0.062	0.050
400	800	800	114.110	113.336	113.978	113.124	0.057	0.051	0.056	0.050
$\alpha = 0.01$										
20	10	10	165.180	137.168	157.031	124.325	0.356	0.052	0.248	0.010
30	15	15	146.182	129.655	142.055	124.257	0.124	0.022	0.088	0.010
40	20	20	139.217	127.413	136.440	124.202	0.068	0.016	0.051	0.010
80	40	40	130.939	125.388	129.758	124.143	0.026	0.012	0.023	0.010
200	100	100	126.672	124.524	126.247	124.089	0.015	0.011	0.014	0.010
400	200	200	125.284	124.222	125.076	124.054	0.012	0.010	0.012	0.010
20	20	20	163.201	138.600	155.800	124.362	0.323	0.059	0.229	0.010
30	30	30	144.864	130.306	141.162	124.181	0.111	0.024	0.081	0.010
40	40	40	138.230	127.811	135.840	124.071	0.062	0.017	0.048	0.010
80	80	80	130.417	125.503	129.422	124.106	0.025	0.013	0.022	0.010
200	200	200	126.534	124.626	126.171	124.130	0.015	0.011	0.014	0.010
400	400	400	125.318	124.373	125.143	124.157	0.012	0.010	0.012	0.010
20	40	40	161.507	139.596	154.745	124.358	0.296	0.065	0.214	0.010
30	60	60	143.690	130.694	140.447	124.204	0.101	0.025	0.076	0.010
40	80	80	137.411	128.090	135.308	124.043	0.057	0.018	0.046	0.010
80	160	160	129.955	125.547	129.090	123.996	0.024	0.012	0.021	0.010
200	400	400	126.457	124.742	126.154	124.248	0.014	0.011	0.014	0.010
400	800	800	125.283	124.433	125.140	124.244	0.012	0.011	0.012	0.010

Note. The closest to α in the values $\alpha_1, \alpha_{\rho_m}, \alpha_\tau$, and α_ϕ of each row is in bold. $\chi_f^2(0.05) = 113.15$, $\chi_f^2(0.01) = 124.12$, $f = 90$.

Table 6: The simulated values for $-2 \log \lambda$, $-2\rho_m \log \lambda$, $-2 \log \tau$, and $-2 \log \phi$ and type I error rates α_1 , α_{ρ_m} , α_τ , and α_ϕ when $(p_1, p_2, p_3) = (5, 5, 5)$ for Case I.

Sample Size			Upper Percentile				Type I Error Rate			
N_1	N_2	N_3	$-2 \log \lambda$	$-2\rho_m \log \lambda$	$-2 \log \tau$	$-2 \log \phi$	α_1	α_{ρ_m}	α_τ	α_ϕ
$\alpha = 0.05$										
20	10	10	242.616	192.408	227.782	163.916	0.918	0.363	0.836	0.055
30	10	10	202.336	172.560	195.157	163.234	0.556	0.121	0.445	0.051
40	10	10	189.893	168.061	185.054	163.246	0.360	0.083	0.286	0.051
80	10	10	175.179	164.303	173.046	163.081	0.152	0.057	0.128	0.050
200	10	10	167.747	163.329	166.928	163.092	0.081	0.051	0.074	0.050
400	10	10	165.398	163.170	164.994	163.135	0.064	0.050	0.061	0.050
20	20	20	239.055	195.081	225.446	163.819	0.898	0.400	0.814	0.054
30	20	20	200.305	173.987	193.915	163.281	0.523	0.136	0.424	0.051
40	20	20	188.566	169.055	184.255	163.200	0.339	0.091	0.273	0.050
80	20	20	174.801	164.752	172.813	163.151	0.147	0.060	0.126	0.050
200	20	20	167.625	163.385	166.845	163.106	0.080	0.052	0.074	0.050
400	20	20	165.369	163.192	164.977	163.114	0.064	0.050	0.061	0.050
20	50	50	235.294	197.421	222.919	163.767	0.873	0.431	0.788	0.054
30	50	50	197.618	175.248	192.105	163.212	0.480	0.150	0.395	0.051
40	50	50	186.604	170.054	182.954	163.130	0.307	0.099	0.253	0.050
80	50	50	173.999	165.330	172.323	163.150	0.139	0.064	0.122	0.050
200	50	50	167.471	163.620	166.757	163.147	0.078	0.053	0.073	0.050
400	50	50	165.324	163.272	164.956	163.123	0.064	0.051	0.061	0.050
$\alpha = 0.01$										
20	10	10	264.515	209.775	247.725	177.093	0.795	0.172	0.652	0.011
30	10	10	218.812	186.612	210.921	176.214	0.315	0.034	0.220	0.010
40	10	10	205.272	181.672	199.989	176.378	0.158	0.020	0.112	0.010
80	10	10	189.289	177.538	186.945	176.127	0.046	0.012	0.036	0.010
200	10	10	181.171	176.399	180.292	176.170	0.019	0.010	0.017	0.010
400	10	10	178.631	176.225	178.196	176.164	0.014	0.010	0.013	0.010
20	20	20	260.578	212.645	245.177	176.900	0.758	0.199	0.620	0.011
30	20	20	216.690	188.219	209.593	176.181	0.286	0.040	0.204	0.010
40	20	20	203.761	182.678	199.033	176.249	0.145	0.023	0.105	0.010
80	20	20	188.738	177.889	186.574	176.102	0.043	0.013	0.035	0.010
200	20	20	181.047	176.468	180.194	176.146	0.019	0.010	0.017	0.010
400	20	20	178.490	176.139	178.041	176.034	0.014	0.010	0.013	0.010
20	50	50	256.731	215.408	242.626	176.916	0.716	0.224	0.585	0.011
30	50	50	213.824	189.619	207.656	176.261	0.249	0.046	0.183	0.010
40	50	50	201.568	183.690	197.568	176.239	0.125	0.025	0.094	0.010
80	50	50	188.076	178.706	186.236	176.334	0.040	0.014	0.033	0.010
200	50	50	180.925	176.765	180.157	176.221	0.018	0.011	0.017	0.010
400	50	50	178.452	176.236	178.048	176.054	0.014	0.010	0.013	0.010

Note. The closest to α in the values $\alpha_1, \alpha_{\rho_m}, \alpha_\tau$, and α_ϕ of each row is in bold. $\chi_f^2(0.05) = 163.12$, $\chi_f^2(0.01) = 176.14$, $f = 135$.

Table 7: The simulated values for $-2 \log \lambda$, $-2\rho_m \log \lambda$, $-2 \log \tau$, and $-2 \log \phi$ and type I error rates α_1 , α_{ρ_m} , α_τ , and α_ϕ when $(p_1, p_2, p_3) = (5, 5, 5)$ for Case II.

Sample Size			Upper Percentile				Type I Error Rate			
N_1	N_2	N_3	$-2 \log \lambda$	$-2\rho_m \log \lambda$	$-2 \log \tau$	$-2 \log \phi$	α_1	α_{ρ_m}	α_τ	α_ϕ
$\alpha = 0.05$										
20	10	10	242.616	192.408	227.782	163.916	0.918	0.363	0.836	0.055
30	15	15	201.122	173.375	194.409	163.246	0.538	0.130	0.434	0.051
40	20	20	188.608	169.092	184.281	163.136	0.338	0.091	0.272	0.050
80	40	40	174.239	165.225	172.477	163.143	0.141	0.063	0.123	0.050
200	100	100	167.259	163.798	166.635	163.146	0.077	0.054	0.072	0.050
400	200	200	165.225	163.516	164.925	163.191	0.063	0.052	0.061	0.050
20	20	20	239.055	195.081	225.446	163.819	0.898	0.400	0.814	0.054
30	30	30	199.032	174.624	193.061	163.248	0.504	0.143	0.411	0.051
40	40	40	187.153	169.939	183.346	163.195	0.315	0.098	0.259	0.050
80	80	80	173.376	165.403	171.863	163.041	0.133	0.064	0.118	0.050
200	200	200	166.925	163.855	166.393	163.052	0.075	0.054	0.071	0.050
400	400	400	164.917	163.400	164.656	163.035	0.061	0.052	0.059	0.050
20	40	40	236.162	197.064	223.569	163.807	0.878	0.426	0.793	0.054
30	60	60	197.168	175.406	191.837	163.213	0.471	0.151	0.389	0.051
40	80	80	185.729	170.354	182.386	163.201	0.293	0.102	0.245	0.050
80	160	160	172.830	165.677	171.527	163.108	0.126	0.066	0.113	0.050
200	400	400	166.831	164.069	166.380	163.223	0.073	0.055	0.070	0.050
400	800	800	164.875	163.510	164.663	163.111	0.060	0.052	0.059	0.050
$\alpha = 0.01$										
20	10	10	264.515	209.775	247.725	177.093	0.795	0.172	0.652	0.011
30	15	15	217.631	187.606	210.218	176.330	0.299	0.038	0.212	0.010
40	20	20	203.742	182.661	198.972	176.053	0.144	0.022	0.104	0.010
80	40	40	187.927	178.205	186.030	176.013	0.041	0.013	0.033	0.010
200	100	100	180.431	176.697	179.763	175.995	0.017	0.011	0.016	0.010
400	200	200	178.325	176.479	177.999	176.143	0.013	0.010	0.013	0.010
20	20	20	260.578	212.645	245.177	176.900	0.758	0.199	0.620	0.011
30	30	30	215.392	188.978	208.755	176.151	0.268	0.043	0.194	0.010
40	40	40	202.166	183.572	198.025	176.270	0.130	0.025	0.097	0.010
80	80	80	187.360	178.744	185.723	176.151	0.038	0.014	0.032	0.010
200	200	200	180.481	177.161	179.902	176.278	0.017	0.012	0.016	0.010
400	400	400	178.182	176.543	177.916	176.135	0.013	0.011	0.013	0.010
20	40	40	257.531	214.895	243.117	176.871	0.725	0.219	0.592	0.011
30	60	60	213.452	189.893	207.507	176.368	0.242	0.046	0.179	0.010
40	80	80	200.798	184.176	197.109	176.153	0.117	0.026	0.090	0.010
80	160	160	186.792	179.061	185.351	176.189	0.035	0.015	0.030	0.010
200	400	400	180.092	177.110	179.606	176.176	0.017	0.011	0.016	0.010
400	800	800	178.308	176.832	178.076	176.346	0.013	0.011	0.013	0.010

Note. The closest to α in the values α_1 , α_{ρ_m} , α_τ , and α_ϕ of each row is in bold. $\chi_f^2(0.05) = 163.12$, $\chi_f^2(0.01) = 176.14$, $f = 135$.

Table 8: The simulated values for $-2 \log \lambda$, $-2\rho_m \log \lambda$, $-2 \log \tau$, and $-2 \log \phi$ and type I error rates $\alpha_1, \alpha_{\rho_m}, \alpha_\tau$, and α_ϕ when $(p_1, p_2, p_3) = (6, 6, 6)$ for Case I.

Sample Size			Upper Percentile				Type I Error Rate			
N_1	N_2	N_3	$-2 \log \lambda$	$-2\rho_m \log \lambda$	$-2 \log \tau$	$-2 \log \phi$	α_1	α_{ρ_m}	α_τ	α_ϕ
$\alpha = 0.05$										
20	10	10	405.478	306.401	372.738	226.217	0.999	0.811	0.995	0.072
30	10	10	291.585	240.920	279.272	222.512	0.828	0.198	0.718	0.052
40	10	10	267.788	231.436	259.856	222.319	0.576	0.110	0.465	0.051
80	10	10	241.924	224.190	238.558	222.072	0.218	0.061	0.179	0.050
200	10	10	229.529	222.391	228.273	222.054	0.095	0.051	0.086	0.050
400	10	10	225.828	222.237	225.210	222.136	0.070	0.051	0.067	0.050
20	20	20	399.318	312.588	368.598	226.168	0.998	0.846	0.993	0.072
30	20	20	288.268	243.546	277.132	222.457	0.799	0.228	0.692	0.052
40	20	20	265.544	233.102	258.402	222.184	0.545	0.124	0.445	0.051
80	20	20	241.300	224.922	238.201	222.090	0.210	0.065	0.174	0.050
200	20	20	229.382	222.531	228.174	222.013	0.095	0.052	0.086	0.050
400	20	20	225.774	222.263	225.169	222.076	0.070	0.051	0.066	0.050
20	50	50	393.397	318.633	364.625	226.157	0.997	0.873	0.991	0.072
30	50	50	283.839	245.902	274.158	222.373	0.756	0.256	0.654	0.051
40	50	50	262.501	235.012	256.411	222.294	0.498	0.140	0.413	0.051
80	50	50	239.941	225.826	237.298	222.106	0.195	0.070	0.166	0.050
200	50	50	229.183	222.961	228.099	222.149	0.093	0.054	0.085	0.050
400	50	50	225.597	222.290	225.027	222.062	0.069	0.051	0.066	0.050
$\alpha = 0.01$										
20	10	10	444.322	335.754	407.065	241.944	0.995	0.638	0.979	0.017
30	10	10	312.219	257.969	298.761	237.655	0.635	0.068	0.483	0.011
40	10	10	286.170	247.323	277.549	237.378	0.329	0.029	0.232	0.010
80	10	10	258.367	239.427	254.736	237.162	0.075	0.013	0.057	0.010
200	10	10	245.192	237.567	243.840	237.232	0.024	0.011	0.021	0.010
400	10	10	241.058	237.225	240.403	237.166	0.015	0.010	0.014	0.010
20	20	20	438.000	342.868	402.734	241.776	0.993	0.692	0.973	0.017
30	20	20	308.767	260.865	296.574	237.745	0.591	0.083	0.452	0.011
40	20	20	283.689	249.030	275.941	237.265	0.301	0.034	0.216	0.010
80	20	20	257.850	240.349	254.492	237.172	0.071	0.014	0.055	0.010
200	20	20	245.099	237.779	243.807	237.189	0.024	0.011	0.021	0.010
400	20	20	240.978	237.231	240.343	237.063	0.015	0.010	0.014	0.010
20	50	50	431.813	349.749	398.429	241.871	0.989	0.734	0.965	0.017
30	50	50	303.916	263.295	293.193	237.491	0.533	0.099	0.411	0.010
40	50	50	280.344	250.987	273.747	237.103	0.262	0.041	0.194	0.010
80	50	50	256.439	241.354	253.565	237.202	0.064	0.016	0.051	0.010
200	50	50	244.635	237.993	243.474	237.133	0.023	0.011	0.020	0.010
400	50	50	240.800	237.270	240.196	236.970	0.015	0.010	0.014	0.010

Note. The closest to α in the values $\alpha_1, \alpha_{\rho_m}, \alpha_\tau$, and α_ϕ of each row is in bold. $\chi_f^2(0.05) = 222.08$, $\chi_f^2(0.01) = 237.15$, $f = 189$.

Table 9: The simulated values for $-2 \log \lambda$, $-2\rho_m \log \lambda$, $-2 \log \tau$, and $-2 \log \phi$ and type I error rates α_1 , α_{ρ_m} , α_τ , and α_ϕ when $(p_1, p_2, p_3) = (6, 6, 6)$ for Case II.

Sample Size			Upper Percentile				Type I Error Rate			
N_1	N_2	N_3	$-2 \log \lambda$	$-2\rho_m \log \lambda$	$-2 \log \tau$	$-2 \log \phi$	α_1	α_{ρ_m}	α_τ	α_ϕ
$\alpha = 0.05$										
20	10	10	405.478	306.401	372.738	226.217	0.999	0.811	0.995	0.072
30	15	15	289.591	242.417	277.970	222.415	0.812	0.216	0.703	0.052
40	20	20	265.610	233.160	258.430	222.211	0.545	0.124	0.444	0.051
80	40	40	240.238	225.563	237.492	222.040	0.198	0.069	0.168	0.050
200	100	100	228.799	223.208	227.829	222.079	0.090	0.056	0.083	0.050
400	200	200	225.385	222.631	224.929	222.139	0.068	0.053	0.065	0.050
20	20	20	399.318	312.588	368.598	226.168	0.998	0.846	0.993	0.072
30	30	30	286.232	244.786	275.783	222.436	0.780	0.243	0.676	0.052
40	40	40	263.145	234.568	256.816	222.150	0.509	0.136	0.421	0.050
80	80	80	239.233	226.243	236.843	222.101	0.186	0.073	0.160	0.050
200	200	200	228.421	223.460	227.608	222.134	0.087	0.057	0.082	0.050
400	400	400	225.109	222.665	224.718	222.039	0.066	0.053	0.064	0.050
20	40	40	394.554	317.428	365.374	226.043	0.998	0.868	0.991	0.071
30	60	60	283.210	246.303	273.709	222.333	0.747	0.260	0.647	0.051
40	80	80	260.978	235.470	255.342	222.223	0.477	0.145	0.398	0.051
80	160	160	238.286	226.642	236.212	222.167	0.175	0.075	0.154	0.050
200	400	400	227.910	223.454	227.212	221.980	0.084	0.057	0.079	0.050
400	800	800	224.995	222.796	224.669	222.079	0.065	0.054	0.063	0.050
$\alpha = 0.01$										
20	10	10	444.322	335.754	407.065	241.944	0.995	0.638	0.979	0.017
30	15	15	310.127	259.608	297.356	237.521	0.609	0.077	0.465	0.010
40	20	20	283.953	249.262	276.140	237.261	0.302	0.035	0.218	0.010
80	40	40	256.799	241.112	253.832	237.241	0.066	0.016	0.052	0.010
200	100	100	244.292	238.323	243.276	237.170	0.022	0.011	0.020	0.010
400	200	200	240.587	237.648	240.100	237.115	0.015	0.011	0.014	0.010
20	20	20	438.000	342.868	402.734	241.776	0.993	0.692	0.973	0.017
30	30	30	306.312	261.959	294.826	237.543	0.563	0.091	0.431	0.011
40	40	40	281.334	250.782	274.396	237.165	0.271	0.040	0.199	0.010
80	80	80	255.449	241.578	252.883	237.186	0.060	0.017	0.049	0.010
200	200	200	243.761	238.467	242.874	236.965	0.021	0.012	0.019	0.010
400	400	400	240.595	237.982	240.176	237.388	0.015	0.011	0.014	0.010
20	40	40	433.129	348.463	399.361	241.846	0.990	0.727	0.967	0.017
30	60	60	303.122	263.620	292.685	237.401	0.523	0.101	0.403	0.010
40	80	80	279.136	251.854	272.958	237.344	0.244	0.043	0.183	0.010
80	160	160	254.460	242.025	252.214	237.208	0.055	0.017	0.045	0.010
200	400	400	243.474	238.714	242.728	237.213	0.020	0.012	0.019	0.010
400	800	800	240.380	238.031	240.028	237.312	0.014	0.011	0.014	0.010

Note. The closest to α in the values $\alpha_1, \alpha_{\rho_m}, \alpha_\tau$, and α_ϕ of each row is in bold. $\chi_f^2(0.05) = 222.08$, $\chi_f^2(0.01) = 237.15$, $f = 189$.

6 Conclusions

We discussed the simultaneous test of the mean vector and the covariance matrix with three-step monotone missing data. We proposed two modified LRT statistics $(-2 \log \tau, -2 \log \phi)$ by decomposing the LR and correcting it by extracting the LR of the simultaneous test and the test of variance for complete data. We also proposed a modified LRT statistic $(-2\rho_m \log \lambda)$ via linear interpolation. From the simulation results, the modified LRT statistic $-2 \log \phi$, which was modified only for the LR part of the test of variance for the complete data, gave the most accurate results. The focus of future works will be the extension to k -step monotone missing data and two-sample problem.

Appendix

To derive the MLEs of the mean vector and the covariance matrix, we consider the following transformation matrix \mathbf{Z} :

$$\mathbf{Z} = \left(\begin{array}{cc|c} \mathbf{I}_{p_1} & \mathbf{O} & \mathbf{O} \\ -\boldsymbol{\Sigma}_{21}\boldsymbol{\Sigma}_{11}^{-1} & \mathbf{I}_{p_2} & \mathbf{O} \\ \hline -\boldsymbol{\Sigma}_{3(12)}\boldsymbol{\Sigma}_{(12)(12)}^{-1} & & \mathbf{I}_{p_3} \end{array} \right).$$

In this case, the transformed vector \mathbf{y}_j is

$$\mathbf{y}_j = \mathbf{Z}\mathbf{x}_j = \left(\begin{array}{c} \mathbf{x}_{1j} \\ -\boldsymbol{\Sigma}_{21}\boldsymbol{\Sigma}_{11}^{-1}\mathbf{x}_{1j} + \mathbf{x}_{2j} \\ -\boldsymbol{\Sigma}_{3(12)}\boldsymbol{\Sigma}_{(12)(12)}^{-1} \begin{pmatrix} \mathbf{x}_{1j} \\ \mathbf{x}_{2j} \end{pmatrix} + \mathbf{x}_{3j} \end{array} \right).$$

The transformed parameters are defined as

$$\boldsymbol{\eta} = \begin{pmatrix} \boldsymbol{\eta}_1 \\ \boldsymbol{\eta}_2 \\ \boldsymbol{\eta}_3 \end{pmatrix} = \begin{pmatrix} \boldsymbol{\mu}_1 \\ -\boldsymbol{\Sigma}_{21}\boldsymbol{\Sigma}_{11}^{-1}\boldsymbol{\mu}_1 + \boldsymbol{\mu}_2 \\ -\boldsymbol{\Sigma}_{3(12)}\boldsymbol{\Sigma}_{(12)(12)}^{-1} \begin{pmatrix} \boldsymbol{\mu}_1 \\ \boldsymbol{\mu}_2 \end{pmatrix} + \boldsymbol{\mu}_3 \end{pmatrix},$$

$$\boldsymbol{\Delta} = \left(\begin{array}{cc|c} \boldsymbol{\Delta}_{11} & \boldsymbol{\Delta}_{12} & \boldsymbol{\Delta}_{13} \\ \boldsymbol{\Delta}_{21} & \boldsymbol{\Delta}_{22} & \boldsymbol{\Delta}_{23} \\ \hline \boldsymbol{\Delta}_{31} & \boldsymbol{\Delta}_{32} & \boldsymbol{\Delta}_{33} \end{array} \right) = \left(\begin{array}{c|c} \boldsymbol{\Delta}_{(12)(12)} & \boldsymbol{\Delta}_{(12)3} \\ \hline \boldsymbol{\Delta}_{3(12)} & \boldsymbol{\Delta}_{33} \end{array} \right),$$

where

$$\begin{aligned} \boldsymbol{\Delta}_{11} &= \boldsymbol{\Sigma}_{11}, \\ \boldsymbol{\Delta}_{12} &= \boldsymbol{\Delta}'_{21} = \boldsymbol{\Sigma}_{11}^{-1}\boldsymbol{\Sigma}_{12}, \\ \boldsymbol{\Delta}_{22} &= \boldsymbol{\Sigma}_{22 \cdot 1} = \boldsymbol{\Sigma}_{22} - \boldsymbol{\Sigma}_{21}\boldsymbol{\Sigma}_{11}^{-1}\boldsymbol{\Sigma}_{12}, \\ \boldsymbol{\Delta}_{(12)3} &= \boldsymbol{\Delta}'_{3(12)} = \boldsymbol{\Sigma}_{(12)(12)}^{-1}\boldsymbol{\Sigma}_{(12)3}, \\ \boldsymbol{\Delta}_{33} &= \boldsymbol{\Sigma}_{33 \cdot 12} = \boldsymbol{\Sigma}_{33} - \boldsymbol{\Sigma}_{3(12)}\boldsymbol{\Sigma}_{(12)(12)}^{-1}\boldsymbol{\Sigma}_{(12)3}. \end{aligned}$$

We note that the pair $(\boldsymbol{\eta}, \boldsymbol{\Delta})$ is in one-to-one correspondence with $(\boldsymbol{\mu}, \boldsymbol{\Sigma})$. In this case, the likelihood function is

$$L(\boldsymbol{\eta}, \boldsymbol{\Delta}) = C |\boldsymbol{\Delta}_{11}|^{-\frac{1}{2}N} |\boldsymbol{\Delta}_{22}|^{-\frac{1}{2}(N_1+N_2)} |\boldsymbol{\Delta}_{33}|^{-\frac{1}{2}N_1} \exp \left\{ -\frac{1}{2} \sum_{j=1}^N (\mathbf{y}_{1j} - \boldsymbol{\eta}_1)' \boldsymbol{\Delta}_{11}^{-1} (\mathbf{y}_{1j} - \boldsymbol{\eta}_1) \right\} \\ \times \exp \left\{ -\frac{1}{2} \sum_{j=1}^{N_1+N_2} (\mathbf{y}_{2j} - \boldsymbol{\eta}_2)' \boldsymbol{\Delta}_{22}^{-1} (\mathbf{y}_{2j} - \boldsymbol{\eta}_2) \right\} \exp \left\{ -\frac{1}{2} \sum_{j=1}^{N_1} (\mathbf{y}_{3j} - \boldsymbol{\eta}_3)' \boldsymbol{\Delta}_{33}^{-1} (\mathbf{y}_{3j} - \boldsymbol{\eta}_3) \right\},$$

where C is a constant. The MLEs can be obtained by differentiating $\log L(\boldsymbol{\eta}, \boldsymbol{\Delta})$ with respect to $\boldsymbol{\eta}$ and $\boldsymbol{\Delta}$, respectively.

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